### Paper Specific Instructions

- 1. The examination is of 3 hours duration. There are a total of 60 questions carrying 100 marks. The entire paper is divided into three sections, **A**, **B** and **C**. All sections are compulsory. Questions in each section are of different types.
- Section A contains a total of 30 Multiple Choice Questions (MCQ). Each MCQ type question has four choices out of which only one choice is the correct answer. Questions Q.1 Q.30 belong to this section and carry a total of 50 marks. Q.1 Q.10 carry 1 mark each and Questions Q.11 Q.30 carry 2 marks each.
- **3.** Section B contains a total of 10 Multiple Select Questions (MSQ). Each MSQ type question is similar to MCQ but with a difference that there may be one or more than one choice(s) that are correct out of the four given choices. The candidate gets full credit if he/she selects all the correct answers only and no wrong answers. Questions Q.31 Q.40 belong to this section and carry 2 marks each with a total of 20 marks.
- **4.** Section C contains a total of 20 Numerical Answer Type (NAT) questions. For these NAT type questions, the answer is a real number which needs to be entered using the virtual keyboard on the monitor. No choices will be shown for these type of questions. Questions Q.41 Q.60 belong to this section and carry a total of 30 marks. Q.41 Q.50 carry 1 mark each and Questions Q.51 Q.60 carry 2 marks each.
- 5. In all sections, questions not attempted will result in zero mark. In Section A (MCQ), wrong answer will result in NEGATIVE marks. For all 1 mark questions, 1/3 marks will be deducted for each wrong answer. For all 2 marks questions, 2/3 marks will be deducted for each wrong answer. In Section B (MSQ), there is NO NEGATIVE and NO PARTIAL marking provisions. There is NO NEGATIVE marking in Section C (NAT) as well.
- **6.** Only Virtual Scientific Calculator is allowed. Charts, graph sheets, tables, cellular phone or other electronic gadgets are **NOT** allowed in the examination hall.
- 7. The Scribble Pad will be provided for rough work.

Special Instructions/Useful Data				
	All angles are in radian			
$\mathbb{R}$	Set of all real numbers			
$\mathbb{R}^{n}$	$\{(x_1, x_2, \dots, x_n) : x_i \in \mathbb{R}, 1 \le i \le n\}$			
$M^T$	Transpose of the matrix M			
f'	Derivative of the function <i>f</i>			
P(E)	Probability of the event <i>E</i>			
E(X)	Expectation of the random variable <i>X</i>			
Var(X)	Variance of the random variable <i>X</i>			
i.i.d.	Independently and identically distributed			
U(a,b)	Continuous uniform distribution on $(a, b)$ , $-\infty < a < b < \infty$			
Φ( <i>a</i> )	$\frac{1}{\sqrt{2\pi}} \int_{-\infty}^{a} e^{-x^2/2} dx$			
Г(р)	The gamma function $\Gamma(p) = \int_0^\infty e^{-t} t^{p-1} dt, \qquad p > 0$			
n!	The factorial function $n! = n \cdot (n-1) \cdots 3 \cdot 2 \cdot 1$			
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## **SECTION – A**

#### **MULTIPLE CHOICE QUESTIONS (MCQ)**

# Q. 1 – Q.10 carry one mark each.

Q.1 Let  $\{a_n\}_{n\geq 1}$  be a sequence of real numbers such that  $a_1 = 2$  and, for  $n \geq 1$ ,

$$a_{n+1} = \frac{2 a_n + 1}{a_n + 1}.$$

Then

- (A)  $1.5 \le a_n \le 2$ , for all natural number  $n \ge 1$
- (B) there exists a natural number  $n \ge 1$  such that  $a_n > 2$
- (C) there exists a natural number  $n \ge 1$  such that  $a_n < 1.5$
- (D) there exists a natural number  $n \ge 1$  such that  $a_n = \frac{1+\sqrt{5}}{2}$

(B)  $e^{-1}$ 

Q.2 The value of

$$\lim_{n\to\infty}\left(1+\frac{2}{n}\right)^{n^2}e^{-2n}$$

(C) e

(D)  $e^{2}$ 

is

(A)  $e^{-2}$ 

- Q.3 Let  $\{a_n\}_{n\geq 1}$  and  $\{b_n\}_{n\geq 1}$  be two convergent sequences of real numbers. For  $n \geq 1$ , define  $u_n = \max\{a_n, b_n\}$  and  $v_n = \min\{a_n, b_n\}$ . Then
  - (A) neither  $\{u_n\}_{n\geq 1}$  nor  $\{v_n\}_{n\geq 1}$  converges
  - (B)  $\{u_n\}_{n\geq 1}$  converges but  $\{v_n\}_{n\geq 1}$  does not converge
  - (C)  $\{u_n\}_{n\geq 1}$  does not converge but  $\{v_n\}_{n\geq 1}$  converges
  - (D) both  $\{u_n\}_{n\geq 1}$  and  $\{v_n\}_{n\geq 1}$  converge

Q.4  
Let 
$$M = \begin{bmatrix} \frac{1}{4} & \frac{3}{4} \\ \frac{3}{5} & \frac{2}{5} \end{bmatrix}$$
. If *I* is the 2 × 2 identity matrix and **0** is the 2 × 2 zero matrix, then  
(A) 20  $M^2 - 13 M + 7 I = \mathbf{0}$   
(B) 20  $M^2 - 13 M - 7 I = \mathbf{0}$   
(C) 20  $M^2 + 13 M + 7 I = \mathbf{0}$   
(D) 20  $M^2 + 13 M - 7 I = \mathbf{0}$ 

Q.5 Let *X* be a random variable with the probability density function

$$f(x) = \begin{cases} \frac{\alpha^p}{\Gamma(p)} e^{-\alpha x} x^{p-1}, & x \ge 0, \alpha > 0, p > 0, \\ 0, & \text{otherwise.} \end{cases}$$

If E(X) = 20 and Var(X) = 10, then  $(\alpha, p)$  is

(A) (2, 20) (B) (2, 40) (C) (4, 20) (D) (4, 40)

Q.6 Let *X* be a random variable with the distribution function

$$F(x) = \begin{cases} 0, & x < 0, \\ \frac{1}{4} + \frac{4x - x^2}{8}, & 0 \le x < 2, \\ 1, & x \ge 2. \end{cases}$$

Then

$$P(X = 0) + P(X = 1.5) + P(X = 2) + P(X \ge 1)$$

equals

(A) 
$$\frac{3}{8}$$
 (B)  $\frac{5}{8}$  (C)  $\frac{7}{8}$  (D) 1

Q.7 Let  $X_1, X_2$  and  $X_3$  be i.i.d. U(0, 1) random variables. Then  $E\left(\frac{X_1+X_2}{X_1+X_2+X_3}\right)$  equals

- (A)  $\frac{1}{3}$  (B)  $\frac{1}{2}$  (C)  $\frac{2}{3}$  (D)  $\frac{3}{4}$
- Q.8 Let  $x_1 = 0, x_2 = 1, x_3 = 2, x_4 = 3$  and  $x_5 = 0$  be the observed values of a random sample of size 5 from a discrete distribution with the probability mass function

$$f(x;\theta) = P(X = x) = \begin{cases} \frac{\theta}{3}, & x = 0, \\ \frac{2\theta}{3}, & x = 1, \\ \frac{1-\theta}{2}, & x = 2, 3 \end{cases}$$

where  $\theta \in [0, 1]$  is the unknown parameter. Then the maximum likelihood estimate of  $\theta$  is

(A)  $\frac{2}{5}$  (B)  $\frac{3}{5}$  (C)  $\frac{5}{7}$  (D)  $\frac{5}{9}$ 

Q.9 Consider four coins labelled as 1, 2, 3 and 4. Suppose that the probability of obtaining a 'head' in a single toss of the  $i^{th}$  coin is  $\frac{i}{4}$ , i = 1, 2, 3, 4. A coin is chosen uniformly at random and flipped. Given that the flip resulted in a 'head', the conditional probability that the coin was labelled either 1 or 2 equals

(A) 
$$\frac{1}{10}$$
 (B)  $\frac{2}{10}$  (C)  $\frac{3}{10}$  (D)  $\frac{4}{10}$ 

Q.10 Consider the linear regression model  $y_i = \beta_0 + \beta_1 x_i + \epsilon_i$ ; i = 1, 2, ..., n, where  $\epsilon_i$ 's are i.i.d. standard normal random variables. Given that

$$\frac{1}{n}\sum_{i=1}^{n} x_{i} = 3.2, \qquad \frac{1}{n}\sum_{i=1}^{n} y_{i} = 4.2, \quad \frac{1}{n}\sum_{j=1}^{n} \left(x_{j} - \frac{1}{n}\sum_{i=1}^{n} x_{i}\right)^{2} = 1.5 \text{ and}$$
$$\frac{1}{n}\sum_{j=1}^{n} \left(x_{j} - \frac{1}{n}\sum_{i=1}^{n} x_{i}\right) \left(y_{j} - \frac{1}{n}\sum_{i=1}^{n} y_{i}\right) = 1.7,$$

the maximum likelihood estimates of  $\beta_0$  and  $\beta_1$ , respectively, are

(A) 
$$\frac{17}{15}$$
 and  $\frac{32}{75}$   
(B)  $\frac{32}{75}$  and  $\frac{17}{15}$   
(C)  $\frac{17}{15}$  and  $\frac{43}{75}$   
(D)  $\frac{43}{75}$  and  $\frac{17}{15}$ 

### Q. 11 – Q. 30 carry two marks each.

Q.11 Let  $f: [-1, 1] \to \mathbb{R}$  be defined by  $f(x) = \frac{x^2 + [\sin \pi x]}{1 + |x|}$ , where [y] denotes the greatest integer less than or equal to y. Then

- (A) f is continuous at  $-\frac{1}{2}$ , 0, 1
- (B) f is discontinuous at  $-1, 0, \frac{1}{2}$
- (C) f is discontinuous at  $-1, -\frac{1}{2}, 0, \frac{1}{2}$
- (D) f is continuous everywhere except at 0
- Q.12 Let  $f, g: \mathbb{R} \to \mathbb{R}$  be defined by  $f(x) = x^2 \frac{\cos x}{2}$  and  $g(x) = \frac{x \sin x}{2}$ . Then
  - (A) f(x) = g(x) for more than two values of x
  - (B)  $f(x) \neq g(x)$ , for all x in  $\mathbb{R}$
  - (C) f(x) = g(x) for exactly one value of x
  - (D) f(x) = g(x) for exactly two values of x
- Q.13 Consider the domain  $D = \{ (x, y) \in \mathbb{R}^2 : x \le y \}$  and the function  $h: D \to \mathbb{R}$  defined by  $h((x, y)) = (x 2)^4 + (y 1)^4, (x, y) \in D.$

Then the minimum value of h on D equals

(A) 
$$\frac{1}{2}$$
 (B)  $\frac{1}{4}$  (C)  $\frac{1}{8}$  (D)  $\frac{1}{16}$ 

Q.14 Let  $M = \begin{bmatrix} X & Y & Z \end{bmatrix}$  be an orthogonal matrix with  $X, Y, Z \in \mathbb{R}^3$  as its column vectors. Then  $Q = X X^T + Y Y^T$ 

- (A) is a skew-symmetric matrix
- (B) is the  $3 \times 3$  identity matrix
- (C) satisfies  $Q^2 = Q$
- (D) satisfies QZ = Z

Q.15 Let  $f: [0,3] \rightarrow \mathbb{R}$  be defined by

$$f(x) = \begin{cases} 0, & 0 \le x < 1, \\ e^{x^2} - e, & 1 \le x < 2, \\ e^{x^2} + 1, & 2 \le x \le 3. \end{cases}$$

Now, define  $F: [0, 3] \rightarrow \mathbb{R}$  by

$$F(0) = 0$$
 and  $F(x) = \int_0^x f(t)dt$ , for  $0 < x \le 3$ .

Then

(A) *F* is differentiable at 
$$x = 1$$
 and  $F'(1) = 0$ 

- (B) *F* is differentiable at x = 2 and F'(2) = 0
- (C) *F* is not differentiable at x = 1
- (D) *F* is differentiable at x = 2 and F'(2) = 1

Q.16 If x, y and z are real numbers such that 4x + 2y + z = 31 and 2x + 4y - z = 19, then the value of 9x + 7y + z

- (A) cannot be computed from the given information
- (B) equals  $\frac{281}{3}$ (C) equals  $\frac{182}{3}$ (D) equals  $\frac{218}{3}$

Q.17 Let 
$$M = \begin{bmatrix} 1 & -1 & 1 \\ 1 & -1 & -1 \end{bmatrix}$$
. If  
 $V = \left\{ (x, y, 0) \in \mathbb{R}^3 : M \begin{bmatrix} x \\ y \\ 0 \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \end{bmatrix} \right\}$  and  $W = \left\{ (x, y, z) \in \mathbb{R}^3 : M \begin{bmatrix} x \\ y \\ z \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \end{bmatrix} \right\}$ ,

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then

- (A) the dimension of V equals 2
- (B) the dimension of W equals 2
- (C) the dimension of V equals 1

(D)  $V \cap W = \{(0,0,0)\}$ 

Q.18 Let *M* be a  $3 \times 3$  non-zero, skew-symmetric real matrix. If *I* is the  $3 \times 3$  identity matrix, then

- (A) M is invertible
- (B) the matrix I + M is invertible
- (C) there exists a non-zero real number  $\alpha$  such that  $\alpha I + M$  is not invertible
- (D) all the eigenvalues of M are real

Q.19 Let *X* be a random variable with the moment generating function

$$M_X(t) = \frac{6}{\pi^2} \sum_{n \ge 1} \frac{e^{t^2/2n}}{n^2}, t \in \mathbb{R}.$$

Then  $P(X \in \mathbb{Q})$ , where  $\mathbb{Q}$  is the set of rational numbers, equals

(A) 0 (B)  $\frac{1}{4}$  (C)  $\frac{1}{2}$  (D)

Q.20 Let X be a discrete random variable with the moment generating function

$$M_X(t) = \frac{(1+3e^t)^2(3+e^t)^3}{1024}, t \in \mathbb{R}$$

Then

(A) 
$$E(X) = \frac{9}{4}$$
  
(B)  $Var(X) = \frac{15}{32}$   
(D)  $P(X = 5) = \frac{3}{1024}$ 

Q.21 Let  $\{X_n\}_{n\geq 1}$  be a sequence of independent random variables with  $X_n$  having the probability density function as

$$f_n(x) = \begin{cases} \frac{1}{2^{n/2}} \Gamma(\frac{n}{2}) \\ 0, \end{cases} e^{-\frac{x}{2}} x^{(\frac{n}{2}-1)}, \quad x > 0, \\ 0, & \text{otherwise.} \end{cases}$$

Then

$$\lim_{n \to \infty} \left[ P\left(X_n > \frac{3}{4}n\right) + P(X_n > n + 2\sqrt{2n}) \right]$$

equals

(A) 
$$1 + \Phi(2)$$
 (B)  $1 - \Phi(2)$  (C)  $\Phi(2)$  (D)  $2 - \Phi(2)$ 

Let X be a Poisson random variable with mean  $\frac{1}{2}$ . Then E((X + 1)!) equals Q.22

(A) 
$$2 e^{-\frac{1}{2}}$$
 (B)  $4 e^{-\frac{1}{2}}$  (C)  $4 e^{-1}$  (D)  $2 e^{-1}$ 

Q.23 Let *X* be a standard normal random variable. Then  $P(X^3 - 2X^2 - X + 2 > 0)$  equals

(A) 
$$2\Phi(1) - 1$$
 (B)  $1 - \Phi(2)$   
(C)  $2\Phi(1) - \Phi(2)$  (D)  $\Phi(2) - \Phi(1)$ 

Q.24 Let X and Y have the joint probability density function

Let X and Y have the joint probability density function  

$$f(x, y) = \begin{cases} 2, & 0 \le x \le y \le 1, \\ 0, & \text{otherwise.} \end{cases}$$
Let  $a = E(Y|X = \frac{1}{2})$  and  $b = Var(Y|X = \frac{1}{2})$ . Then  $(a, b)$  is  
(A)  $\left(\frac{3}{4}, \frac{7}{12}\right)$ 
(B)  $\left(\frac{1}{4}, \frac{1}{48}\right)$ 
(C)  $\left(\frac{1}{4}, \frac{7}{12}\right)$ 
(D)  $\left(\frac{3}{4}, \frac{1}{48}\right)$ 

# Q.25 Let X and Y have the joint probability mass function

$$P(X = m, Y = n) = \begin{cases} \frac{m+n}{21}, & m = 1, 2, 3; n = 1, 2, \\ 0, & \text{otherwise.} \end{cases}$$

Then P(X = 2|Y = 2) equals

(A) 
$$\frac{1}{3}$$
 (B)  $\frac{2}{3}$  (C)  $\frac{1}{2}$  (D)  $\frac{1}{4}$ 

Let X and  $\overline{Y}$  be two independent standard normal random variables. Then the probability density Q.26 function of  $Z = \frac{|X|}{|Y|}$  is

(A) 
$$f(z) = \begin{cases} \frac{\sqrt{1/2}}{\sqrt{\pi}} e^{-\frac{z}{2}} z^{-\frac{1}{2}}, & z > 0, \\ 0, & \text{otherwise} \end{cases}$$
 (B)  $f(z) = \begin{cases} \frac{2}{\sqrt{2\pi}} e^{-z^2/2}, & z > 0, \\ 0, & \text{otherwise} \end{cases}$   
(C)  $f(z) = \begin{cases} e^{-z}, & z > 0, \\ 0, & \text{otherwise} \end{cases}$  (D)  $f(z) = \begin{cases} \frac{2}{\pi} \frac{1}{(1+z^2)}, & z > 0, \\ 0, & \text{otherwise} \end{cases}$ 

Q.27 Let *X* and *Y* have the joint probability density function

$$f(x,y) = \begin{cases} e^{-y}, & 0 < x < y < \infty, \\ 0, & \text{otherwise.} \end{cases}$$

Then the correlation coefficient between *X* and *Y* equals

(A) 
$$\frac{1}{3}$$
 (B)  $\frac{1}{\sqrt{3}}$  (C)  $\frac{1}{\sqrt{2}}$  (D)  $\frac{2}{\sqrt{3}}$ 

Q.28 Let  $x_1 = -2$ ,  $x_2 = 1$  and  $x_3 = -1$  be the observed values of a random sample of size three from a discrete distribution with the probability mass function

$$f(x;\theta) = P(X = x) = \begin{cases} \frac{1}{2\theta + 1}, & x \in \{-\theta, -\theta + 1, \dots, 0, \dots, \theta\}, \\ 0, & \text{otherwise,} \end{cases}$$

where  $\theta \in \Theta = \{1, 2, ...\}$  is the unknown parameter. Then the method of moment estimate of  $\theta$  is

(A) 1 (B) 2 (C) 3 (D) 4

Q.29 Let X be a random sample from a discrete distribution with the probability mass function

$$f(x;\theta) = P(X = x) = \begin{cases} \frac{1}{\theta}, & x = 1, 2, ..., \theta \\ 0, & \text{otherwise,} \end{cases}$$

where  $\theta \in \Theta = \{20, 40\}$  is the unknown parameter. Consider testing

$$H_0: \theta = 40$$
 against  $H_1: \theta = 20$ 

at a level of significance  $\alpha = 0.1$ . Then the uniformly most powerful test rejects  $H_0$  if and only if

- (A)  $X \le 4$  (B) X > 4 

   (C)  $X \ge 3$  (D) X < 3
- Q.30 Let  $X_1$  and  $X_2$  be a random sample of size 2 from a discrete distribution with the probability mass function

$$f(x;\theta) = P(X=x) = \begin{cases} \theta, & x=0, \\ 1-\theta, & x=1, \end{cases}$$

where  $\theta \in \Theta = \{0.2, 0.4\}$  is the unknown parameter. For testing  $H_0: \theta = 0.2$  against  $H_1: \theta = 0.4$ , consider a test with the critical region

$$C = \{ (x_1, x_2) \in \{0, 1\} \times \{0, 1\} : x_1 + x_2 < 2 \}.$$

Let  $\alpha$  and  $\beta$  denote the probability of Type I error and power of the test, respectively. Then  $(\alpha, \beta)$  is

(A) (0.36, 0.74)	(B) (0.64, 0.36)
(C) (0.05, 0.64)	(D) (0.36, 0.64)

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# **SECTION - B**

#### **MULTIPLE SELECT QUESTIONS (MSQ)**

### Q. 31 – Q. 40 carry two marks each.

Q.31 Let  $\{a_n\}_{n\geq 1}$  be a sequence of real numbers such that

$$a_n = \sum_{k=n+1}^{2n} \frac{1}{k}, \quad n \ge 1.$$

Then which of the following statement(s) is (are) true?

- (A)  $\{a_n\}_{n\geq 1}$  is an increasing sequence
- (B)  $\{a_n\}_{n\geq 1}$  is bounded below
- (C)  $\{a_n\}_{n\geq 1}$  is bounded above
- (D)  $\{a_n\}_{n\geq 1}$  is a convergent sequence
- Q.32 Let  $\sum_{n \ge 1} a_n$  be a convergent series of positive real numbers. Then which of the following statement(s) is (are) true?
  - (A)  $\sum_{n\geq 1} (a_n)^2$  is always convergent
  - (B)  $\sum_{n\geq 1} \sqrt{a_n}$  is always convergent
  - (C)  $\sum_{n \ge 1} \frac{\sqrt{a_n}}{n}$  is always convergent
  - (D)  $\sum_{n\geq 1} \frac{\sqrt{a_n}}{n^{1/4}}$  is always convergent
- Q.33 Let  $\{a_n\}_{n\geq 1}$  be a sequence of real numbers such that  $a_1 = 3$  and, for  $n \geq 1$ ,

$$a_{n+1} = \frac{a_n^2 - 2 a_n + 4}{2}.$$

Then which of the following statement(s) is (are) true?

- (A)  $\{a_n\}_{n\geq 1}$  is a monotone sequence
- (B)  $\{a_n\}_{n\geq 1}$  is a bounded sequence
- (C)  $\{a_n\}_{n\geq 1}$  does not have finite limit, as  $n \to \infty$
- (D)  $\lim_{n \to \infty} a_n = 2$

Q.34 Let  $f: \mathbb{R} \to \mathbb{R}$  be defined by

$$f(x) = \begin{cases} x^4(2 + \sin\frac{1}{x}), & x \neq 0, \\ 0, & x = 0. \end{cases}$$

Then which of the following statement(s) is (are) true?

- (A) f attains its minimum at 0
- (B) f is monotone
- (C) f is differentiable at 0
- (D)  $f(x) > 2x^4 + x^3$ , for all x > 0
- Q.35 Let *P* be a probability function that assigns the same weight to each of the points of the sample space  $\Omega = \{1,2,3,4\}$ . Consider the events  $E = \{1,2\}, F = \{1,3\}$  and  $G = \{3,4\}$ . Then which of the following statement(s) is (are) true?
  - (A) E and F are independent
  - (B) E and G are independent
  - (C) F and G are independent
  - (D) E, F and G are independent
- Q.36 Let  $X_1, X_2, ..., X_n, n \ge 5$ , be a random sample from a distribution with the probability density function

$$f(x;\theta) = \begin{cases} e^{-(x-\theta)}, & x \ge \theta, \\ 0, & \text{otherwise} \end{cases}$$

where  $\theta \in \mathbb{R}$  is the unknown parameter. Then which of the following statement(s) is (are) true?

- (A) A 95% confidence interval of  $\theta$  has to be of finite length
- (B)  $(\min\{X_1, X_2, ..., X_n\} + \frac{1}{n}\ln(0.05), \min\{X_1, X_2, ..., X_n\})$  is a 95% confidence interval of  $\theta$
- (C) A 95% confidence interval of  $\theta$  can be of length 1
- (D) A 95% confidence interval of  $\theta$  can be of length 2
- Q.37 Let  $X_1, X_2, ..., X_n$  be a random sample from  $U(0, \theta)$ , where  $\theta > 0$  is the unknown parameter. Let  $X_{(n)} = \max\{X_1, X_2, ..., X_n\}$ . Then which of the following is (are) consistent estimator(s) of  $\theta^3$ ?
  - (A)  $8 X_n^3$  (B)  $X_{(n)}^3$ (C)  $\left(\frac{2}{n} \sum_{i=5}^n X_i\right)^3$  (D)  $\frac{n X_{(n)}^3 + 1}{n+1}$

Q.38 Let  $X_1, X_2, ..., X_n$  be a random sample from a distribution with the probability density function

$$f(x;\theta) = \begin{cases} c(\theta) \ e^{-(x-\theta)}, & x \ge 2\theta, \\ 0, & \text{otherwise,} \end{cases}$$

where  $\theta \in \mathbb{R}$  is the unknown parameter. Then which of the following statement(s) is (are) true?

- (A) The maximum likelihood estimator of  $\theta$  is  $\frac{\min\{X_1, X_2, ..., X_n\}}{2}$
- (B)  $c(\theta) = 1$ , for all  $\theta \in \mathbb{R}$
- (C) The maximum likelihood estimator of  $\theta$  is min{ $X_1, X_2, ..., X_n$ }
- (D) The maximum likelihood estimator of  $\theta$  does not exist

Q.39 Let  $X_1, X_2, ..., X_n$  be a random sample from a distribution with the probability density function

$$f(x;\theta) = \begin{cases} \theta^2 x \ e^{-\theta x}, & x > 0, \\ 0, & \text{otherwise,} \end{cases}$$

where  $\theta > 0$  is the unknown parameter. If  $Y = \sum_{i=1}^{n} X_i$ , then which of the following statement(s) is (are) true?

- (A) *Y* is a complete sufficient statistic for  $\theta$
- (B)  $\frac{2n}{N}$  is the uniformly minimum variance unbiased estimator of  $\theta$
- (C)  $\frac{2n-1}{v}$  is the uniformly minimum variance unbiased estimator of  $\theta$
- (D)  $\frac{2n+1}{v}$  is the uniformly minimum variance unbiased estimator of  $\theta$
- Q.40 Let  $X_1, X_2, ..., X_n$  be a random sample from  $U(\theta, \theta + 1)$ , where  $\theta \in \mathbb{R}$  is the unknown parameter. Let  $U = \max\{X_1, X_2, ..., X_n\}$  and  $V = \min\{X_1, X_2, ..., X_n\}$ . Then which of the following statement(s) is (are) true?
  - (A) U is a consistent estimator of  $\theta$
  - (B) V is a consistent estimator of  $\theta$
  - (C) 2U V 2 is a consistent estimator of  $\theta$
  - (D) 2V U + 1 is a consistent estimator of  $\theta$

#### **SECTION - C**

#### NUMERICAL ANSWER TYPE (NAT)

# Q. 41 – Q. 50 carry one mark each.

Q.41 Let  $\{a_n\}_{n\geq 1}$  be a sequence of real numbers such that

 $a_n = \frac{1+3+5+\dots+(2n-1)}{n!}, \ n \ge 1.$ 

Then  $\sum_{n \ge 1} a_n$  converges to \_\_\_\_\_

Q.42 Let

$$S = \Big\{ (x,y) \in \mathbb{R}^2 \colon x,y \geq 0,$$

$$\sqrt{4 - (x - 2)^2} \le y \le \sqrt{9 - (x - 3)^2}$$

Then the area of *S* equals \_\_\_\_\_

Q.43 Let  $S = \{(x, y) \in \mathbb{R}^2 : |x| + |y| \le 1\}$ . Then the area of S equals \_\_\_\_\_

Q.44 Let

$$J = \frac{1}{\pi} \int_0^1 t^{-\frac{1}{2}} (1-t)^{\frac{3}{2}} dt$$

Then the value of *J* equals

- Q.45 A fair die is rolled three times independently. Given that 6 appeared at least once, the conditional probability that 6 appeared exactly twice equals \_\_\_\_\_\_
- Q.46 Let *X* and *Y* be two positive integer valued random variables with the joint probability mass function

$$P(X = m, Y = n) = \begin{cases} g(m) \ h(n), & m, n \ge 1, \\ 0, & \text{otherwise,} \end{cases}$$
  
where  $g(m) = \left(\frac{1}{2}\right)^{m-1}$ ,  $m \ge 1$  and  $h(n) = \left(\frac{1}{3}\right)^n$ ,  $n \ge 1$ . Then  $E(X Y)$  equals \_\_\_\_\_\_

Q.47 Let *E*, *F* and *G* be three events such that

$$P(E \cap F \cap G) = 0.1, P(G|F) = 0.3 \text{ and } P(E|F \cap G) = P(E|F)$$

Then  $P(G|E \cap F)$  equals \_\_\_\_\_

Q.48 Let  $A_1, A_2$  and  $A_3$  be three events such that

$$P(A_i) = \frac{1}{3}, i = 1, 2, 3; P(A_i \cap A_j) = \frac{1}{6}, 1 \le i \ne j \le 3 \text{ and } P(A_1 \cap A_2 \cap A_3) = \frac{1}{6}.$$

Then the probability that none of the events  $A_1, A_2, A_3$  occur equals \_\_\_\_\_

Q.49 Let  $X_1, X_2, ..., X_n$  be a random sample from the distribution with the probability density function

$$f(x) = \frac{1}{4} e^{-|x-4|} + \frac{1}{4} e^{-|x-6|}, x \in \mathbb{R}.$$

Then  $\frac{1}{n}\sum_{i=1}^{n} X_i$  converges in probability to \_\_\_\_\_

Q.50 Let  $x_1 = 1.1, x_2 = 2.2$  and  $x_3 = 3.3$  be the observed values of a random sample of size three from a distribution with the probability density function

$$f(x;\theta) = \begin{cases} \frac{1}{\theta} e^{-x/\theta}, & x > 0, \\ 0, & \text{otherwise,} \end{cases}$$

where  $\theta \in \Theta = \{1, 2, ...\}$  is the unknown parameter. Then the maximum likelihood estimate of  $\theta$  equals \_\_\_\_\_\_

### Q. 51 – Q. 60 carry two marks each.

Q.51 Let  $f: \mathbb{R} \to \mathbb{R}$  be a differentiable function such that f' is continuous on  $\mathbb{R}$  with f'(3) = 18. Define

$$g_n(x) = n\left(f\left(x+\frac{5}{n}\right) - f\left(x-\frac{2}{n}\right)\right).$$

Then  $\lim_{n\to\infty} g_n(3)$  equals \_\_\_\_\_

- Q.52 Let  $M = \sum_{i=1}^{4} X_i X_i^T$ , where  $X_1^T = \begin{bmatrix} 1 & -1 & 1 & 0 \end{bmatrix}$ ,  $X_2^T = \begin{bmatrix} 1 & 1 & 0 & 1 \end{bmatrix}$ ,  $X_3^T = \begin{bmatrix} 1 & 3 & 1 & 0 \end{bmatrix}$  and  $X_4^T = \begin{bmatrix} 1 & 1 & 1 & 0 \end{bmatrix}$ . Then the rank of M equals \_\_\_\_\_\_
- Q.53 Let  $f: \mathbb{R} \to \mathbb{R}$  be a differentiable function with  $\lim_{x \to \infty} f(x) = \infty$  and  $\lim_{x \to \infty} f'(x) = 2$ . Then

$$\lim_{x\to\infty} \left(1 + \frac{f(x)}{x^2}\right)^x$$

equals \_\_\_\_\_

Q.54 The value of

equals

e of  $\int_0^{\frac{\pi}{2}} \left( \int_0^x e^{\sin y} \sin x \, dy \right) \, dx$ 

Q.55 Let *X* be a random variable with the probability density function

$$f(x) = \begin{cases} 4 x^{\kappa}, & 0 < x < 1, \\ x - \frac{x^2}{2}, & 1 \le x < 2, \\ 0, & \text{otherwise,} \end{cases}$$

where *k* is a positive integer. Then  $P\left(\frac{1}{2} < X < \frac{3}{2}\right)$  equals \_\_\_\_\_

Q.56 Let X and Y be two discrete random variables with the joint moment generating function

$$M_{X,Y}(t_1, t_2) = \left(\frac{1}{3} e^{t_1} + \frac{2}{3}\right)^2 \left(\frac{2}{3} e^{t_2} + \frac{1}{3}\right)^3, t_1, t_2 \in \mathbb{R}$$

Then P(2X + 3Y > 1) equals \_\_\_\_\_

Q.57 Let  $X_1, X_2, X_3$  and  $X_4$  be i.i.d. discrete random variables with the probability mass function

$$P(X_1 = n) = \begin{cases} \frac{3^{n-1}}{4^n}, & n = 1, 2, ..., \\ 0, & \text{otherwise.} \end{cases}$$

Then  $P(X_1 + X_2 + X_3 + X_4 = 6)$  equals \_\_\_\_\_

Q.58 Let *X* be a random variable with the probability mass function

$$P(X = n) = \begin{cases} \frac{1}{10}, & n = 1, 2, \cdots, 10, \\ 0, & \text{otherwise.} \end{cases}$$

Then  $E(\max\{X, 5\})$  equals \_\_\_\_\_

Q.59 Let *X* be a sample observation from  $U(\theta, \theta^2)$  distribution, where  $\theta \in \Theta = \{2,3\}$  is the unknown parameter. For testing

 $H_0: \theta = 2$  against  $H_1: \theta = 3$ ,

let  $\alpha$  and  $\beta$  be the size and power, respectively, of the test that rejects  $H_0$  if and only if  $X \ge 3.5$ . Then  $\alpha + \beta$  equals \_\_\_\_\_\_

Q.60 A fair die is rolled four times independently. For i = 1, 2, 3, 4, define

 $Y_i = \begin{cases} 1, & \text{if 6 appears in the } i^{th} \text{ throw,} \\ 0, & \text{otherwise.} \end{cases}$ 

Then  $P(\max\{Y_1, Y_2, Y_3, Y_4\} = 1)$  equals \_\_\_\_\_

# END OF THE QUESTION PAPER

Paper Code : MS				
Q No.	Question Type (QT)	Section	Key/Range (KY)	
1	MCQ	А	А	
2	MCQ	А	А	
3	MCQ	А	D	
4	MCQ	А	В	
5	MCQ	А	В	
6	MCQ	А	С	
7	MCQ	А	С	
8	MCQ	А	В	
9	MCQ	А	С	
10	MCQ	А	D	
11	MCQ	А	В	
12	MCQ	А	D	
13	MCQ	А	С	
14	MCQ	A	с	
15	MCQ	A	A	
16	МСQ	A	D	
17	MCQ	А	С	
18	MCQ	А	В	
19	MCQ	А	А	
20	MCQ	А	А	
21	MCQ	А	D	
22	MCQ	А	В	
23	MCQ	А	С	

Paper Code : MS				
Q No.	Question Type (QT)	Section	Key/Range (KY)	
24	MCQ	А	D	
25	MCQ	А	А	
26	MCQ	А	D	
27	MCQ	А	С	
28	MCQ	А	В	
29	MCQ	А	А	
30	MCQ	А	D	
31	MSQ	В	A,B,C,D	
32	MSQ	В	A,C	
33	MSQ	В	A,C	
34	MSQ	В	A,C	
35	MSQ	В	A,C	
36	MSQ	В	B,C,D	
37	MSQ	В	B,C,D	
38	MSQ	В	А	
39	MSQ	В	A,C	
40	MSQ	В	B,C,D	
41	NAT	С	5.40 to 5.50	
42	NAT	с	7.80 to 7.90	
43	NAT	с	1.90 to 2.10	
44	NAT	с	0.35 to 0.40	
45	NAT	с	0.16 to 0.17	
46	NAT	С	2.50 to 3.50	

Paper Code : MS							
Q No.	Question Type (QT)	Section	Key/Range (KY)				
47	NAT	с	0.25 to 0.35				
48	48 NAT		0.30 to 0.40				
49	NAT	С	4.75 to 5.25				
50	<b>50</b> NAT		1.90 to 2.10				
51	<b>51</b> NAT		125 to 127				
52	52 NAT		3.0 to 3.0				
53	53 NAT		7.25 to 7.50				
54	54 NAT		1.70 to 1.80				
55	NAT	С	0.85 to 0.90				
56	NAT	С	0.97 to 0.99				
57	<b>57</b> NAT		0.01 to 0.03				
58	58 NAT		6.25 to 6.75				
59	NAT	С	1.10 to 1.20				
60	NAT	C	0.50 to 0.53				
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